ON THE SUMS OF DIGITS OF FIBONACCI NUMBERS

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1. INTRODUCTION

The problem of determining which integers k are equal to the sum of the digits of F_k was first brought to my attention at the Fibonacci Conference in Pullman, Washington, this summer (1994). Professor Dan Fielder presented this as an open problem, having obtained all solutions for $k \leq 2000$. There seemed to be fairly many solutions in base 10, and it was not clear whether there were infinitely many. Shortly after hearing the problem, it occurred to me why there were so many solutions. If one assumes that the digits F_k are independently uniformly randomly distributed, then one expects S(k), the sum of the digits of F_k , to be approximately $\frac{9}{2}k\log_{10}\alpha$, where $\alpha = \frac{1}{2}(1+\sqrt{5}) \approx 1.61803$ is the golden mean. Since $\frac{9}{2}\log_{10}\alpha \approx 0.94044$, we expect $S(k) \approx$ 0.94044k. Since this is close to k, we expect many solutions to S(k) = k, at least for reasonably small k. However, as k gets large, we expect S(k)/k to deviate from 0.94044 by less and less. Thus, it appears that, for some integer n_0 , the ratio S(k)/k never gets as large as 1 for $k > n_0$, so S(k) = k has no solutions for $k > n_0$, and thus has finitely many solutions. In this paper, I present two closely related probabilistic models to predict the number of solutions. More generally, they predict N(b, n), the number of solutions to S(k, b) = k for $k \le n$, where S(k, b) is the sum of the digits of F_k in base b [thus, S(k; 10) = S(k)]. Let N(b) denote the total number of solutions in base b [thus, $N(b; n) \rightarrow N(b)$ as $n \rightarrow \infty$]. Both models predict finite values of N(b) for each base b. In the simpler model, N(10) is estimated to be 18.24 ± 3.86 , compared with the actual value N(10; 20000) = 20.

2. THE NAIVE MODEL

In this model I assume that the digits of F_k are independently uniformly randomly distributed among $\{0, 1, ..., b-1\}$ for each positive integer k and each fixed base $b \ge 4$. [It is fairly easy to prove that the only solutions to S(k; b) = k are 0 and 1 when b = 2 or 3. The proof involves showing that, for all sufficiently large k, we have $(b-1)(1+\log_b F_k) < k$.] Now let $\alpha = \frac{1}{2}(1+\sqrt{5})$ and $\beta = \frac{1}{2}(1-\sqrt{5})$. Then

$$F_k = \frac{\alpha^k - \beta^k}{\sqrt{5}} = \frac{\alpha^k}{\sqrt{5}} + o(1) \text{ for } k \to \infty.$$

The number of digits of F_k in base *b* is approximately the base-*b* logarithm of this number, $k \log_b \alpha - \log_b \sqrt{5} \approx k \log_b \alpha = k\gamma$, where $\gamma = \log_b \alpha$ and I neglect terms of order 1. In this model, the expected value of each digit of F_k is $\frac{1}{2}(b-1)$ and the standard deviation (SD) is $\sqrt{\frac{1}{12}(b^2-1)}$ (see [2], pp. 80-86). Therefore, the expected value of S(k; b) is approximately $\overline{S} = \frac{k}{2}(b-1)\gamma$ and the SD is approximately $\sigma = \sqrt{\frac{k}{12}(b^2-1)\gamma}$. Let $\mathcal{P}_1(k; \ell)$ denote the probability that $\widetilde{S}(k; b) = \ell$, where $\widetilde{S}(k; b)$ is distributed as the sum $Y_{k,1} + Y_{k,2} + \cdots + Y_{k, [k\gamma]}$, the $Y_{k,j}$ being independent random variables, each uniformly distributed over $\{0, 1, \dots, b-1\}$. According to the

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central limit theorem ([2], pp. 165-77), if k is reasonably large, the probability distribution is approximately Gaussian, so

$$\mathcal{P}_1(k;\ell) \approx \frac{1}{\sigma\sqrt{2\pi}} \exp\left[\frac{-(\overline{S}-\ell)^2}{2\sigma^2}\right] = \sqrt{\frac{6}{k\pi\gamma(b^2-1)}} \exp\left[\frac{-6\left(k\gamma(\frac{b-1}{2})-\ell\right)^2}{k\gamma(b^2-1)}\right]$$

Let $\mathcal{P}_1(k) = \mathcal{P}_1(k; k)$; this is the estimated percentage of Fibonacci numbers $F_{k'}$ for k' near k whose base-b digits sum to the index k'. We have $\mathcal{P}_1(k) \approx Ae^{-Bk} / \sqrt{k}$, where

$$A = \sqrt{\frac{6}{\pi \gamma (b^2 - 1)}} \quad \text{and} \quad B = \frac{6(\gamma (\frac{b-1}{2}) - 1)^2}{\gamma (b^2 - 1)}$$

Incidentally, it is clear that the only solutions k for which $F_k < b$ are those for which $F_k = k$, namely 0, 1, and possibly 5 (if b > 5). We might as well put in these solutions by hand. Thus, in the model, we only calculate $\mathcal{P}_1(k)$ for k for which $F_k \ge b$ and add N_0 to the final result upon summing the probabilities, where $N_0 = 3$ if b > 5, otherwise $N_0 = 2$. Thus, our estimate for N(b; n) in this model is

$$N_1(b;n) = N_0 + \sum_{\substack{k \le n \\ F_k \ge b}} \mathcal{P}_1(k) \approx N_0 + \sum_{\substack{k \le n \\ F_k \ge b}} \frac{Ae^{-Bk}}{\sqrt{k}}$$

and the standard deviation of this estimate is [assuming that the S(k, b) are uncorrelated for different values of k]

$$\Delta_1(b;n) = \sqrt{\sum_{\substack{k \le n \\ F_k \ge b}} \mathcal{P}_1(k)(1-\mathcal{P}_1(k))} \approx \sqrt{\sum_{\substack{k \le n \\ F_k \ge b}} \frac{Ae^{-Bk}(\sqrt{k}-Ae^{-Bk})}{k}$$

This model gives good results for some bases, but not all. The next model is an improvement which seems to yield accurate results for all bases.

3. THE IMPROVED MODEL

In this model, I still assume that the digits of F_n are uniformly distributed over $\{0, 1, ..., b-1\}$, but with one restriction, namely, their sum modulo b-1. It is well known that the sum of the base-10 digits of a number a is congruent to $a \mod 9$. In general, the same applies to the sum of the digits in base $b \mod b-1$. Thus, we have the restriction $S(k; b) \equiv F_k \pmod{b-1}$. In particular, k cannot be a solution to S(k; b) = k unless $k \equiv F_k \pmod{b-1}$. This latter equation is not too difficult to solve. Upon solving it, we end up with a restriction of the form

$$k \mod q \in S. \tag{1}$$

Here, q = [b-1, p], where p = per(b-1) is the period of the Fibonacci sequence modulo b-1 and S is a specified subset of $\{0, 1, ..., q-1\}$. If k does not satisfy the above condition, it need not be considered, since the sum of its digits cannot equal F_n . On the other hand, if k does satisfy the condition, we know that the sum of its digits is congruent to F_n modulo b-1. In the improved model, we take this restriction into account and otherwise assume a uniform random distribution of digits in F_n . In analogy to $\tilde{S}(k; b)$, let $\hat{S}(k; b)$ be distributed as the sum $Y_{k,1} + \cdots + Y_{k,\lfloor k\gamma \rfloor}$, the $Y_{k,j}$ being random variables uniformly distributed over 0, 1, ..., b-1 and independent except for the restriction that $Y_{k,1} + \cdots + Y_{k,\lfloor k\gamma \rfloor} \equiv F_k \pmod{b-1}$. We now estimate the probability $\mathcal{P}_2(k)$ that $\hat{S}(k; b) = k$ to be b-1 times our earlier estimate in the case where k satisfies (1) and zero otherwise, i.e.,

$$\mathcal{P}_{2}(k) = \begin{cases} (b-1)\mathcal{P}_{1}(k) & k \mod q \in S, \\ 0 & k \mod q \notin S. \end{cases}$$

Thus, in this model, the expectation and SD of N(b; n) are approximately

$$N_2(b;n) = N_0 + \sum_{\substack{k \le n \\ F_k \ge b}} \mathcal{P}_2(k) \approx N_0 + (b-1) \sum_{\substack{k \le n \\ F_k \ge b \\ k \mod q \in S}} \frac{Ae^{-bk}}{\sqrt{k}}$$

and

$$\Delta_2(b;n) = \sqrt{\sum_{\substack{k \le n \\ F_k \ge b}} \mathcal{P}_2(k)(1-\mathcal{P}_2(k))} \approx \sqrt{(b-1)\sum_{\substack{k \le n \\ F_k \ge b \\ k \bmod q \in S}} \frac{Ae^{-Bk}(\sqrt{k}-Ae^{-Bk})}{k}$$

As an example of how to calculate S, consider b = 8. In this case, p = per(7) = 16 and q = [7, 16] = 112. To determine S, we first tabulate k mod 16 and $F_k \mod 7$ for each congruence class of k mod 16. Next, below the line, we tabulate the unique solutions modulo 112 to the congruences $x \equiv k \pmod{16}$ and $x \equiv F_k \pmod{7}$. Since (16, 7) = 1, by the Chinese Remainder Theorem, each of these solutions exists and is unique.

<i>k</i> mod 16	0	1	2	3	4	5	6	7	8	9	10	11	12	13	14	15	
$F_k \mod 7$	0	1	1	2	3	5	1	6	0	6	6	5	4	2	6	1	
<i>k</i> mod 112	0	1	50	51	52	5	22	55	56	41	90	75	60	93	62	15	

Thus, $S = \{0, 1, 5, 15, 22, 41, 50, 51, 52, 55, 56, 60, 62, 75, 90, 93\}$. Note that, in this example, the pair of congruences $k \equiv j \pmod{p}$ and $k \equiv F_j \mod b - 1$ has a solution mod q for every integer $j \mod p$. This is because, in this example, b - 1 = 7 and p = 16 are coprime. In general, this is not the case. For example, for b = 10, we get p = 24, which is not coprime to b - 1 = 9. Thus, if we constructed a similar table for b = 10, we would expect to get some simultaneous congruences without solutions. This is in fact the case, i.e., the pair of congruences $k \equiv 2 \pmod{24}$ and $k \equiv F_2 = 1 \pmod{9}$ has no solutions. We expect only one-third (eight) of them to have a solution, since (9, 24) = 3. In fact, we do get eight. For b = 10, we find $S = \{0, 1, 5, 10, 31, 35, 36, 62\}$ and q = 72.

One might wonder by about how much $N_1(k; b)$ and $N_2(k; b)$ differ. To first order, they differ by a multiplicative factor depending on b, i.e., $N_2(b; n) \approx M(b)N_1(b; n)$. Recall that in going from the first model to the second, we selected s out of every q congruence classes modulo q, where s = #S. Also, we multiplied the corresponding probabilities by b-1. Thus, M(b) = (b-1)s/q. For some bases, M(b) = 1, so the predictions of both models are essentially the same. This is true in particular whenever b-1 and p are coprime, and also in some other cases, like b = 10. However, there are other bases for which $M(b) \neq 1$; in fact, the difference can be quite

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large! For instance, for b = 11, we find p = q = 60 and s = 14, hence $M(11) = 10 \times 14/60 = 7/3$, which is greater than 2. Thus, for b = 11, the second model predicts over twice as many solutions to S(k; 11) = k as the first model. In this case, as we will see, the second model agrees well with the known data; the first does not.

4. COMPARISON OF MODELS WITH "EXPERIMENT"

Every good scientist knows that the best way to test a model or theory is to see how well its predictions agree with experimental data. In this case, my "experiment" was a computer program I wrote and ran on my Macintosh LCII to determine S(k; b) given $k \le 20000$ and $b \le 20$. Incidentally, it is not necessary to calculate the Fibonacci numbers directly, only to store the digits in an array. Also, only two Fibonacci arrays need to be stored at one time. Nevertheless, trying to compute for k > 20000 presented memory problems, at least for the method I used. Still, this turned out to be sufficient for determining with high certainty all solutions to S(k; b) = kexcept for b = 11.

Here I present all the solutions I found for $4 \le b \le 20$ and $k \le n$.

b=4, n=1000:	0	1						
b=5, n=1000:	0	1						
b=6, n=1000:	0	1	5	9	15	35		
b=7, n=1000:	0	1	5	7	11	12	53	
b=8, n=1000:	0	1	5	22	41			
b=9, n=5000:	0	1	5	29	77	149	312	
b=10, n=20000:	0 175 540	1 180 946	5 216 1188	10 251 2222	31 252	35 360	62 494	72 504
b=11, n=20000:	0 61	1	5	13	41 185	53 193	55 215	60 265
	269 617	353 629	355 630	385 653	397 713	437	481	493 780
	889	905	960	1013	1025	1045	1205	1320
	1405	1435	1913	1981	2125	2153	2280	2297
	2389 2610	2413 2633	2460 2730	2465	2509 2845	2893	2549 2915	2609 3041
	3055 3721	3155 3749	3209 3757	3360 3761	3475	3485 3865	3521 3929	3641 3941
	4075 5489	4273 5490	4301 5700	4650 5917	4937 6169	$\begin{array}{c} 5195 \\ 6253 \\ \end{array}$	5209 6335	$\begin{array}{c} 5435\\ 6361 \end{array}$
	6373 7349	$\begin{array}{c} 6401 \\ 7577 \end{array}$	$\begin{array}{c} 6581 \\ 7595 \end{array}$	$\begin{array}{c} 6593 \\ 7693 \end{array}$	$\begin{array}{c} 6701 \\ 7740 \end{array}$	6750 7805	$6941 \\ 7873$	7021 8009
	$\begin{array}{c} 8017\\9120\end{array}$	$8215 \\ 9133$	$8341 \\ 9181$	$8495 \\ 9269$	8737 9277	$\frac{8861}{9535}$	$8970 \\ 9541$	8995 9737
	$9935 \\ 12029$	$9953 \\ 12175$	$10297 \\ 12353$	$10609 \\ 12461$	$10789 \\ 12565$	$10855 \\ 12805$	$\frac{11317}{12893}$	$11809 \\ 13855$
	$14381 \\ 16177$	$14550 \\ 16789$	$14935 \\ 16837$	$15055 \\ 17065$	$15115 \\ 17237$	$15289 \\ 17605$	15637 17681	15709 17873
	17941 18990	17993 19135	$18193 \\ 19140$	18257 19375	$18421 \\ 19453$	18515 19657	18733 19873	18865

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b=12, n=20000:	0	1	5	13	14	89	96	123
	221	387	419	550	648	749	866	892
	1105	2037						
b=13, n=5000:	0	1	5	12	24	25	36	48
	53	72	73	132	156	173	197	437
	444	485	696	769	773			
b=14, n=3000:	0	1	5	8	11	27	34	181
	192	194						
b=15, n=2000:	0	1	5					
h = 16 n = 2000	0	1	5	10	60	101		
0 = 10, n = 2000.	0	1	0	10	00	101		
b=17, n=1000:	0	1	5					
b=18, n=1000:	0	1	5	60				
b=19, n=1000:	0	1	5	31	36			
b=20, n=1000:	0	1	5	21	22			

Next, I tabulated $N_1(b; n) \pm \Delta_1(b; n)$, $N_2(b; n) \pm \Delta_2(b; n)$, N(b; n), $(\frac{b-1}{2})\log_b \alpha$, and M(b) for the above pairs (b, n). Note how N(b; n) increases as $(\frac{b-1}{2})\log_b \alpha$ approaches 1.

b	n	$N_1(b;n)\pm \Delta_1(b;n)$	$N_2(b;n)\pm\Delta_2(b;n)$	N(b;n)	$(\frac{b-1}{2})\log_b \alpha$	M(b)
4	1000	2.25 ± 0.49	2.43 ± 0.61	2	0.52068	1.00
5	1000	2.67 ± 0.79	2.76 ± 0.72	2	0.59799	1.00
6	1000	4.17 ± 1.05	5.04 ± 1.25	6	0.67142	2.00
7	1000	5.10 ± 1.41	6.18 ± 1.42	7	0.74188	1.50
8	1000	6.61 ± 1.85	5.84 ± 1.47	5	0.80995	1.00
9	5000	9.40 ± 2.48	8.57 ± 2.09	7	0.87604	1.00
10	20000	18.24 ± 3.86	17.77 ± 3.46	20	0.94044	1.00
11	20000	79.71 ± 8.72	180.95 ± 12.82	183	1.00340	2.33
12	20000	17.03 ± 3.71	17.01 ± 3.28	18	1.06510	1.00
13	5000	9.71 ± 2.56	15.73 ± 3.08	21	1.12566	2.00
14	3000	7.15 ± 2.01	8.22 ± 1.62	10	1.18522	1.00
15	2000	5.93 ± 1.69	4.70 ± 1.19	3	1.24387	1.00
16	2000	5.21 ± 1.47	7.16 ± 1.62	6	1.30170	2.00
17	1000	4.75 ± 1.31	3.94 ± 0.90	3	1.35877	1.00
18	1000	4.42 ± 1.18	4.69 ± 1.06	4	1.41515	1.00
19	1000	4.10 ± 1.08	4.12 ± 0.95	5	1.47088	1.50
20	1000	4.01 ± 1.00	4.54 ± 0.97	5	1.52601	1.00

As one can see, the first model does not make accurate predictions for each base. In particular, its predictions for bases 11 and 13 are off by roughly 12 and 4.5 standard deviations, respectively. On the other hand, the second model seems to agree well with the known data for each base. For 12 out of 17 bases, its predictions are correct within one SD, and all 17 predictions are correct within two SD's. (The largest deviation, found for b = 13, is -1.71 SD's.) Furthermore, there does not seem to be a directional bias of the model. Eight out of 17 of the predicted values are too high; the other 9 are too low. Thus, the second model looks good.

5. PREDICTING THE UNKNOWN

With this in mind, we can use the second model to make predictions for which we are unable to calculate at present. In particular, we can estimate N(11), the total number of solutions to S(k, 11) = k in base 11, as well as the value of the largest one. We can also estimate the probability that we missed some solutions in each of the other bases we looked at. For these bases, I was careful to calculate out to large enough n so that these probabilities should be very small.

I calculated $N_2(11; n) \pm \Delta_2(11; n)$ for $200000 \le n \le 4000000$ in intervals of 200000. Here are the results:

	n	$N_2($	(11; n)	$\pm \Delta_{2}$	$_{2}(11; n)$
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200000	490.38 ± 21.70
400000	595.89 ± 24.00
600000	641.02 ± 24.93
800000	662.32 ± 25.35
1000000	672.83 ± 25.56
1200000	678.16 ± 25.66
1400000	680.91 ± 25.71
1600000	682.34 ± 25.74
1800000	683.10 ± 25.76
2000000	683.50 ± 25.77
2200000	683.72 ± 25.77
2400000	683.83 ± 25.77
2600000	683.89 ± 25.77
2800000	683.93 ± 25.77
3000000	683.94 ± 25.77
3200000	683.95 ± 25.77
3400000	683.96 ± 25.77
3600000	683.96 ± 25.77
3800000	683.96 ± 25.77
4000000	683.97 ± 25.77

As can be seen, the results converge rapidly for large n. Let N'(b; n) denote the estimated number of solutions to S(k; b) = k for k > n. Then we have

$$N'(b; n) = \sum_{\substack{k > n \\ k \bmod q \in S}} \frac{Ae^{-Bk}}{\sqrt{k}} \approx M(b) \sum_{k > n} \frac{Ae^{-Bk}}{\sqrt{k}} \approx \frac{M(b)A}{\sqrt{B}} \int_{Bn}^{\infty} \frac{e^{-x}dx}{\sqrt{x}}$$
$$\approx M(b)A \sqrt{\frac{\pi}{B}} \operatorname{erfc} \sqrt{Bn} \approx \frac{M(b)A}{B\sqrt{n}} e^{-Bn},$$

where I make the change of variables $y = \sqrt{x}$ in the integral to get the error function term. In the last step, I use an asymptotic expansion of erfc [1].

I next tabulated N'(b,n) for the pairs (b,n) used, except for b=11, where I used n=14000000, the largest n for which I have estimated N(b, n). Since N' is much less than 1 in each case, the values of N' listed are the approximate probabilities that there is a solution to S(k; b) = k for k > n. I also tabulated the corresponding values of A, B, and M(b). Note that for every base less than 20, except 11, N'(b; n) is less than 10^{-6} ; in fact, the sums of these entries is roughly 10⁻⁶. Thus, if this model is accurate, there is about one chance in a million that I have missed any solutions in these bases. Also, note that the table of estimates of $N_2(11; n)$ can be used to estimate the largest solution to S(k; 11) = k.

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b	n	Α	В	M(b)	N'(b;n)
4	1000	0.606	2.65×10^{-1}	1.000	7.6×10^{-117}
5	1000	0.516	1.35×10^{-1}	1.000	2.5×10^{-60}
6	1000	0.451	6.89×10^{-2}	2.000	4.7×10^{-31}
7	1000	0.401	$3.37 imes 10^{-2}$	1.500	1.3×10^{-15}
8	1000	0.362	1.49×10^{-2}	1.000	2.7×10^{-7}
9	5000	0.330	$5.26 imes 10^{-3}$	1.000	$2.3 imes 10^{-12}$
10	20000	0.304	1.03×10^{-3}	1.000	2.4×10^{-9}
11	4000000	0.282	2.89×10^{-6}	2.333	1.1×10^{-3}
12	20000	0.263	9.18×10^{-4}	1.000	2.1×10^{-9}
13	5000	0.246	3.01×10^{-3}	2.000	6.8×10^{-7}
14	3000	0.232	$5.79 imes 10^{-3}$	1.000	1.6×10^{-8}
15	2000	0.219	8.97×10^{-3}	1.000	7.2×10^{-9}
16	2000	0.208	1.23×10^{-2}	2.000	1.7×10^{-11}
17	1000	0.198	1.58×10^{-2}	1.000	5.5×10^{-8}
18	1000	0.188	1.92×10^{-2}	1.000	1.4×10^{-9}
19	1000	0.180	2.26×10^{-2}	2.000	5.2×10^{-11}
20	1000	0.172	2.59×10^{-2}	1.000	1.1×10^{-12}

Suppose one wishes to find *n* such that there is a 50% chance that there are no solutions larger than *n*. According to Poisson statistics, this happens when the $N_2(11; n) = \ln 2 \approx 0.69$. By interpolating in the previous table, we see that this occurs when $n \approx 1.9 \times 10^6$; this is roughly the value we can expect for the largest solution. Calculating S(k; 11) for k up to 2.8×10^6 yields a 96% probability of finding all the solutions, and going up to 4×10^6 yields a 99.9% probability of finding them all. Perhaps someone will do this calculation in the near future.

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