THE NUMBER OF DERANGEMENTS OF A SEQUENCE WITH GIVEN SPECIFICATION

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SECTION 1

Consider sequences

$$\sigma = (\alpha_1, \alpha_2, \ldots, \alpha_N), \qquad (1.1)$$

where $a_j \in \mathbb{Z}_k = \{1, 2, \ldots, k\}$. The sequence is said to have *specification* $[n_1, n_2, \ldots, n_k]$, where the n_j are non-negative integers, $\mathbb{N} = n_1 + n_2 + \cdots + n_k$, if each element j, $1 \le j \le k$, occurs in σ exactly n_j times. The sequence is called a *derangement* provided no element is in a position occupied by it in the sequence

$$(1, 1, \ldots, 1, 2, 2, \ldots, 2, \ldots, k, k, \ldots, k).$$
 (1.2)

Let $P(n_1, n_2, ..., n_k)$ denote the number of possible derangements. Even and Gil is [1] (see also Jackson [2]) have proved the following result.

$$P(n_1, n_2, \ldots, n_k) = (-1)^{n_1 + n_2 + \cdots + n_k} \cdot \int_0^\infty e^{-x} \left\{ \prod_{j=1}^k L_{n_j}(x) \right\} dx, \qquad (1.3)$$

where $L_n(x)$ is the Laguerre polynomial defined by

$$L_n(x) = \sum_{j=0}^{n} (-1)^j \binom{n}{j} \frac{x^j}{j!}.$$
 (1.4)

The object of the present note is to give a simple proof of (1.3) along the lines of the standard proof of the case $n_1 = n_2 = \cdots = n_k = 1$ [3, p. 59]. We also prove some related results.

SECTION 2

Let
$$P(n, m) = P(n_1, ..., n_k; m_1, ..., m_k),$$
 (2.1)

where $0 \le m_j \le n_j$, denote the number of sequences (1.1) in which, for each j, exactly m_j of the values remain in their original position in (1.2). It follows at once from the definition that

$$P(\boldsymbol{n}, \boldsymbol{m}) = P(\boldsymbol{n} - \boldsymbol{m}, \boldsymbol{0}) \prod_{j=1}^{k} {n_j \choose m_j} = P(\boldsymbol{n} - \boldsymbol{m}) \prod_{j=1}^{k} {n_j \choose m_j}, \qquad (2.2)$$

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where $P(\mathbf{n}) = P(n_1, n_2, \ldots, n_k)$. Clearly

$$\sum_{m=0}^{n} P(n, m) = (n_1, n_2, \dots, n_k) = \frac{N!}{n_1! n_2! \dots n_k!},$$

where

$$\sum_{m=0}^{n} \equiv \sum_{m_1=0}^{n_1} \sum_{m_2=0}^{n_2} \cdots \sum_{m_k=0}^{n_k} .$$

Thus, by (2.2),

$$\sum_{m=0}^{n} \binom{n_1}{m_1} \cdots \binom{n_k}{m_k} P(m) = (n_1, n_2, \dots, n_k).$$

This relation is equivalent to

$$P(n) = \sum_{m=0}^{n} (-1)^{N-M} \binom{n_1}{m_1} \dots \binom{n_k}{m_k} (m_1, \dots, m_k)$$

$$= \sum_{m=0}^{n} (-1)^{N-M} \binom{n_1}{m_1} \dots \binom{n_k}{m_k} \frac{M!}{m_1! \dots m_k!},$$
(2.3)

where $M = m_1 + \cdots + m_k$.

SECTION 3

To verify that (2.3) is in agreement with (1.3), we take

$$\int_{0}^{\infty} e^{-x} \left\{ \prod_{j=1}^{k} L_{n_{j}}(x) \right\} dx = \int_{0}^{\infty} e^{-x} \left\{ \prod_{j=1}^{k} \sum_{m=0}^{n_{j}} (-1)^{m_{j}} \binom{n_{j}}{m_{j}} \frac{x^{m_{j}}}{m_{j}!} \right\} dx$$

$$= \sum_{m=0}^{n} (-1)^{M} \binom{n_{1}}{m_{1}!} \cdots \binom{n_{k}}{m_{k}} \frac{1}{m_{1}!} \cdots \frac{1}{m_{k}!} \int_{0}^{\infty} e^{-x} x^{M} dx$$

$$= \sum_{m=0}^{n} (-1)^{M} \binom{n_{1}}{m_{1}!} \cdots \binom{n_{k}}{m_{k}} \frac{M!}{m_{1}! \cdots m_{k}!}.$$

This evidently proves the equivalence of (1.3) and (2.3).

SECTION 4

Put

$$P_{k}(N) = \sum_{n_{1} + \cdots + n_{k} = N} P(n). \tag{4.1}$$

Thus P(n) denotes the total number of derangements from Z_k of length N. Then by (2.3) we have

$$P_{k}(n) = \sum_{n_{1} + \cdots + n_{k} = N} \sum_{m=0}^{n} (-1)^{N-M} \binom{n_{1}}{m_{1}} \cdots \binom{n_{k}}{m_{k}} \frac{M!}{m_{1}! \cdots m_{k}!}$$

$$= \sum_{m_1 + \cdots + m_k = N} (-1)^{N-M} \frac{M!}{m_1! \cdots m_k!} \sum_{n_1 + \cdots + n_k = N} {n_1 \choose m_1} \cdots {n_k \choose m_k},$$

where as above $M = m_1 + \cdots + m_k$. Since the inner sum on the extreme right is equal to

$$\binom{N+k-1}{M+k-1},$$

we get

$$P_{k}(N) = \sum_{m_{1} + \dots + m_{k} \leq N} (-1)^{N-M} \frac{M!}{m_{1}! \dots m_{k}!} \binom{N+k-1}{M+k-1}$$

$$= \sum_{M=0}^{N} (-1)^{N-M} \binom{N+k-1}{M+k-1} \sum_{m_{1} + \dots + m_{k} = M} \frac{M!}{m_{1}! \dots m_{k}!}.$$

By the multinomial theorem

$$\sum_{m_1+\cdots+m_k=M} \frac{M!}{m_1! \cdots m_k!} = k^M,$$

so that

$$P_{k}(N) = \sum_{M=0}^{N} (-1)^{N-M} {N+k-1 \choose M+k-1} k^{M}$$
(4.2)

It follows from (4.2) that

$$k^{k-1}P_{k}(N) = \sum_{m=k-1}^{N+k-1} (-1)^{N+k-m-1} {N+k-m-1 \choose m} k^{m}$$

$$= \sum_{m=0}^{N+k-1} (-1)^{N+k-m-1} {N+k-m-1 \choose m} k^{m} - \sum_{j=0}^{k-2} (-1)^{N+k-j-1} {N+k-j-1 \choose j} k^{j}$$

and therefore

$$P_{k}(N) = k^{1-k} \left\{ (k-1)^{N+k-1} - \sum_{j=0}^{k-2} (-1)^{N+k-j-1} {N+k-j-1 \choose j} k^{j} \right\} \qquad (k \ge 1).$$
 (4.3)

It follows from (4.3) that, for fixed k > 2,

$$P_k(N) \sim k^{1-k}(k-1)^{N+k-1} \qquad (N \to \infty).$$
 (4.4)

On the other hand, if N is fixed and $k \to \infty$, it is evident from (4.2) that

$$P_{k}(N) = \sum_{M=0}^{N} (-1)^{M} {N+k-1 \choose M} k^{N-M} \sim \sum_{M=0}^{N} (-1)^{M} \frac{k^{M}}{M!} k^{N-M},$$

so that

$$P_{k}(N) \sim k^{N} \sum_{k=0}^{N} \frac{(-1)^{M}}{M!} \qquad (k \to \infty).$$
 (4.5)

SECTION 5

Fairly simple generating functions are implied by (4.2). We have first

$$\sum_{N=0}^{\infty} x^{N} \sum_{M=0}^{N} (-1)^{N-M} \binom{N+k-1}{M+k-1} k^{M} = \sum_{M=0}^{\infty} k^{M} x^{M} \sum_{N=0}^{\infty} (-1)^{N} \binom{N+M+k-1}{M+k-1} x^{N}$$

$$= \sum_{M=0}^{\infty} k^{M} x^{M} (1+x)^{-M-k}$$

$$= (1+x)^{-k} \left(1 - \frac{kx}{1+x}\right)^{-1}.$$

Hence

$$\sum_{N=0}^{\infty} P_k(N) x^N = (1+x)^{-k+1} (1+x-kx)^{-1}.$$
 (5.1)

In the next place

$$\begin{split} \sum_{N=0}^{\infty} P_k(N) \frac{x^N}{(N+k-1)!} &= \sum_{N=0}^{\infty} \frac{x^N}{(N+k-1)!} \sum_{M=0}^{N} (-1)^{N-M} \binom{N+k-1}{M+k-1} k^M \\ &= \sum_{M=0}^{\infty} \frac{k^M x^M}{(M+k-1)!} \sum_{N=0}^{\infty} (-1)^N \frac{x^N}{N!} \\ &= e^{-x} \sum_{N=0}^{\infty} \frac{k^M x^M}{(M+k-1)!} \,. \end{split}$$

Thus

$$\sum_{N=0}^{\infty} P_k(N) \frac{x^N}{(N+k-1)!} = (kx)^{-k+1} e^{-x} \left\{ e^{kx} - \sum_{m=0}^{k-2} \frac{k^m x^m}{m} \right\} \qquad (k \ge 1).$$
 (5.2)

It is easily seen that (4.3) is implied by (5.2).

REFERENCES

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